# Ensuring valid inference for Cox hazard ratios after variable selection



Kelly Van Lancker Joint work with

Stijn Vansteelandt and Oliver Dukes



### My perspective on estimands, identification and estimation



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### 1 [Inference for the conditional hazard ratio](#page-2-0)

2 [Can we go model-free?](#page-30-0)





## Conditional versus unconditional estimands



## Randomized trials with time-to-event endpoints

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 $\Box$  Usually pre-specified.

### **Censoring**

- Logrank test: censoring is (statistically) independent of survival time
- $\blacksquare$  Cox model adjusting for L: censoring is independent of survival time, given treatment A and baseline covariate L

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	- 0 Model misspecification

■ Variable selection procedures can help in choosing a model (with the right variables)!

### Data-adaptive methods

#### ■ Consider the hazard function for the Cox PH model

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\lambda\{t|A\} = \lambda_0'(t)\exp\{\alpha_0 A\}
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 $\blacksquare$  One common strategy: adjust for L iff significantly associated with outcome, conditional on exposure (e.g., at the 5% level)



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## **Censoring**

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- Suppose  $L$  has a moderate effect on outcome, but a strong effect on censoring
- Because censoring implies information loss and may even reduce variation in L in the risk set, variable selection in the outcome model will rarely pick up L when fitting

 $\lambda \{t | A, L\} = \lambda_0(t) \exp \{ \alpha A + \beta L \}.$ 

Upon removing  $L$  from the model, bias induced by informative censoring can result in highly inflated Type I errors

 $\blacksquare$  Our "hesitation" whether or not to adjust for  $L$  translates into a complex mixture distribution of the test statistic

$$
\tilde{Z}_{\alpha} = \begin{cases} Z_{\alpha} & \text{if adjusted for } L \\ Z_{\alpha_0} & \text{if not adjusted for } L, \end{cases}
$$



### Impact of Variable Selection



Results from the test statistic jumping back and forth between  $Z_{\alpha}$  and  $Z_{\alpha_0}$  $\Box$  distribution of the latter might not be centered at zero

This creates bias and inefficiency, and invalidates standard  $\sim$ inference.

#### Results obtained by post-Lasso



 $n = 400; p = 30; A \stackrel{d}{=} Ber(0.5); L \stackrel{d}{=} N(0, \mathbb{I})$  $\mathcal{T} \stackrel{d}{=} \exp(\lambda_{\mathcal{T}})$ , with  $\lambda_{\mathcal{T}} = \exp(b \cdot \nu_{\mathcal{T}} L)$  and  $\nu_{\mathcal{T}} = (1, 1/2, \ldots, 1/9, 1/10, 0_{11}, \ldots, 0_{30})'$  $C \stackrel{d}{=} \exp(\lambda_C)$ , with  $\lambda_C = \exp(\gamma_1 \cdot A + g \cdot \nu_C L)$  and  $\nu_c = (1, 1/2, \ldots, 1/5, 1, 1/2, \ldots, 1/5, 0_{11}, \ldots, 0_{30})'$ 

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**Perform inference on**  $\alpha_{\mu}$  **by conventional methods, based on** robust SE (obtained via standard statistical software).

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## The proportional hazards assumption

### ■ Hazards have been argued to be non-proportional in many settings.

Figure. Nonproportional Hazards and Survival Curves in 3 Hypothetical Trials Comparing a Treatment vs a Control



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Figure, Nonproportional Hazards and Survival Curves in 3 Hypothetical Trials Comparing a Treatment vs a Control



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#### What are we estimating in the Cox model when the proportional hazards assumption fails?

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- $\Box$  When the model is wrong:
	- No good understanding of what the partial likelihood estimator converges to.
	- What we infer depends on the estimator we use.
	- $\blacksquare$  The target of the standard estimator depends on the censoring distribution. (Struthers and Kalbfleisch, 1986; Whitney et al., 2019)

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■ This highlights the benefits of choosing an estimand in a model-free way.

- $\Box$  The estimand may coincide with the model parameter when assumptions hold...
- $\Box$  ...but otherwise still captures the scientific question.

(van der Laan and Rose, 2011; Vansteelandt and Dukes, 2020)

Reconsider the model

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\lambda(t|A,L) = \lambda_0(t) \exp\{\alpha A + \beta L\}
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#### **Such estimands now exist.**

 $\Box$  Estimation methods allow for flexible machine learning methods.

(Whitney et al., 2019; Vansteelandt et al. 2022)

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## Summary

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- We often think of conditional causal effects as parameters in regression models.
- So long as we specify our estimand in advance, we have some freedom in letting the data choose our model, whilst maintaining type I error/interval coverage.
- Our estimand could be a regression parameter, or (even better) defined in a model-free way.
- $\blacksquare$  The latter ensures that always return something that answers the question of interest.

## Thank you for your attention!

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The opinions in this presentation are of the author and do not necessarily represent those of anyone else.

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